

Jeffreys' Prior and Invariance

Theorem 3.1 (Jeffreys' prior invariance) + a worked example

(Lecture slides)

Motivation: “Objective” priors and re-parameterisation

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- Desideratum: prior beliefs should not change just because we re-label the parameter.
- Jeffreys' idea: use the **Fisher information** as a canonical notion of “how sensitive” the likelihood is to changes in the parameter.

Change of variables for priors (recall)

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- This is the standard Jacobian rule for transforming random variables.
- We'll use it to formalise what “invariance” means for priors.

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- The expectation is taken w.r.t. $Y \sim \pi(\cdot | \theta)$.
- Intuition: large $I(\theta)$ means the log-likelihood curves sharply near θ .

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For a scalar parameter θ , define the prior

$$\pi_J(\theta) \propto \sqrt{I(\theta)}.$$

Then for any one-to-one differentiable re-parameterisation $\psi = g(\theta)$, the induced prior on ψ satisfies

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- “Same prior information” in any coordinate system.
- This is why Jeffreys' prior is often called an **objective** (or parameterisation-invariant) prior.

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If this holds, then taking square roots gives

$$\sqrt{I(\psi)} = \sqrt{I(\theta)} \left| \frac{d\theta}{d\psi} \right|,$$

which will match the Jacobian rule for densities and prove invariance.

Proof sketch (Step 1): chain rule for the score

Start from the log-likelihood as a function of ψ :

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Here $\frac{\partial}{\partial \theta} \log \pi(Y | \theta)$ is the **score function**.

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Product rule + chain rule

$$\frac{\partial^2}{\partial \psi^2} \ell(\psi) = \underbrace{\frac{\partial^2}{\partial \theta^2} \log \pi(Y | \theta) \left(\frac{d\theta}{d\psi} \right)^2}_{(A)} + \underbrace{\frac{\partial}{\partial \theta} \log \pi(Y | \theta) \cdot \frac{d^2 \theta}{d\psi^2}}_{(B)}.$$

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Insert (A)+(B):

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$$\mathbb{E} \left[\frac{\partial}{\partial \theta} \log \pi(Y | \theta) \right] = 0.$$

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Therefore

$$I(\psi) = \left(\frac{d\theta}{d\psi} \right)^2 \left(-\mathbb{E} \left[\frac{\partial^2}{\partial \theta^2} \log \pi(Y | \theta) \right] \right) = I(\theta) \left(\frac{d\theta}{d\psi} \right)^2$$

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Transform it to ψ using the Jacobian rule:

$$\pi_J(\psi) = \pi_J(\theta(\psi)) \left| \frac{d\theta}{d\psi} \right| \propto \sqrt{I(\theta)} \left| \frac{d\theta}{d\psi} \right| = \sqrt{I(\psi)}.$$

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Conclusion

Jeffreys' prior is invariant under one-to-one re-parameterisations.

How to compute Jeffreys' prior in practice

Recipe

Given likelihood $\pi(y | \theta)$:

- 1 Write the log-likelihood $\ell(\theta) = \log \pi(y | \theta)$.
- 2 Compute $\frac{\partial^2}{\partial \theta^2} \ell(\theta)$.
- 3 Take expectation w.r.t. $Y \sim \pi(\cdot | \theta)$:

$$I(\theta) = -\mathbb{E} \left[\frac{\partial^2}{\partial \theta^2} \ell(\theta) \right].$$

- 4 Set Jeffreys prior: $\pi_J(\theta) \propto \sqrt{I(\theta)}$.

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Often, we only need $I(\theta)$ up to *proportionality*.

Example: Binomial model (bots on a platform)

Model (Example 3.1 revisited)

$$Y \mid \theta \sim \text{Bin}(n, \theta), \quad \theta \in (0, 1).$$

Likelihood:

$$\pi(y \mid \theta) = \binom{n}{y} \theta^y (1 - \theta)^{n-y}.$$

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Second derivative:

$$\ell''(\theta) = -\frac{y}{\theta^2} - \frac{n - y}{(1 - \theta)^2}.$$

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By definition,

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Therefore

$$I(\theta) = \frac{n}{\theta(1 - \theta)}.$$

Example: Jeffreys prior and identification

Jeffreys prior:

$$\pi_J(\theta) \propto \sqrt{I(\theta)} = \sqrt{\frac{n}{\theta(1-\theta)}} \propto \frac{1}{\sqrt{\theta(1-\theta)}}.$$

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So the **invariant** prior for the binomial probability is $\text{Beta}(1/2, 1/2)$.

Interpretation: what does $\text{Beta}(1/2, 1/2)$ look like?

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- “Objective” does not mean “flat”: uniform prior $\text{Beta}(1, 1)$ is *not* invariant under re-parameterisation.

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- Binomial example: $\pi_J(\theta) \propto \theta^{-1/2}(1 - \theta)^{-1/2}$, so $\theta \sim \text{Beta}(1/2, 1/2)$.

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Takeaway

Jeffreys' prior is not magic; it is a principled default driven by invariance.

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(Optional bridge) Why we care: towards large-sample agreement

- Jeffreys: “choose priors invariantly” (objective defaults).
- Next big theorem (Bernstein–von Mises): with lots of data, posterior \approx Normal around the MLE (under conditions).
- So Bayesian and frequentist answers often agree asymptotically.

Questions?